

# Three-parametric Marcenko-Pastur Density

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## 1. Original and two-parametric Marcenko-Pastur density

Consider  $M \times N$  ( $M \geq N$ ) random matrices  $K = (K_{jk})$  such that the elements are complex, i.i.d and normally distributed with zero mean and variance 1. This setting is described as

$$\Re K_{jk} \sim N(0, 1/2), \quad \Im K_{jk} \sim N(0, 1/2), \quad j = 1, \dots, M, \quad k = 1, \dots, N. \quad (1.1)$$

We consider a statistical ensemble of  $N \times N$  Hermitian random matrices  $L$  defined by

$$L = K^\dagger K. \quad (1.2)$$

We denote the eigenvalues of  $L$  as  $X_j$ ,  $j = 1, \dots, N$ . In the scaling limit,  $N \rightarrow \infty$ ,  $M \rightarrow \infty$  with  $N/M = r \in (0, 1]$ , the empirical distribution of  $\{X_j/M\}$  converges to a deterministic measure. The limit measure has a finite support in  $\mathbb{R}$  and it is explicitly given as a function of the parameter  $r$ .

$$\rho(x; r) = \frac{\sqrt{(x-x_L(x; r))(x_R(x; r)-x)}}{2\pi r x} \mathbf{1}_{(x_L(r), x_R(r))}(x), \quad x_L(r) := (1 - \sqrt{r})^2, \quad x_R(r) := (1 + \sqrt{r})^2. \quad (1.3)$$

A dynamical extension of eigenvalue distribution of Wishart-matrix ensemble is realized by the solution of the following system of stochastic differential equations (SDEs),

$$dX_j^N(t) = 2 \sqrt{X_j^N(t)} dB_j(t) + 2(\nu + 1)dt + 4X_j^N(t) \sum_{\substack{1 \leq k \leq N \\ k \neq j}} \frac{1}{X_j^N(t) - X_k^N(t)}, \quad j = 1, \dots, N, \quad t \geq 0, \quad (1.4)$$

where  $\nu = M - N$  and  $B_j(t)$ ,  $t \geq 0$ ,  $1 \leq j \leq N$  are independent one-dimensional standard Brownian motions. For  $\rho_\xi(x; r, t)$ ,  $t \geq 0$  with initial distribution  $\xi$ , we define the Green function (the resolvent)  $G_\xi(z; r, t)$  by the Stieltjes transform of  $\rho_\xi$ . Then we can prove that this solves the following nonlinear partial differential equation (PDE),

$$\frac{\partial G_\xi}{\partial t} = -\frac{\partial G_\xi}{\partial z} + r \left\{ \frac{\partial G_\xi}{\partial z} - 2z G_\xi \frac{\partial G_\xi}{\partial z} - G_\xi^2 \right\}, \quad t \geq 0. \quad (1.5)$$

Under the initial condition that all particles are concentrated on the origin, Green function  $G_{\delta_0}(z; r, t)$  is given by the solution of the equation,

$$z = \frac{1}{G_{\delta_0}(z)} + \frac{t}{1 - r t G_{\delta_0}(z)}, \quad z \in \mathbb{C} \setminus \mathbb{R}, \quad r \in (0, 1], \quad t \geq 0. \quad (1.6)$$

The probability density on the time-dependent extension is obtained by solving the equation and using the Sokhotski-Plemelj theorem [1],

$$\rho(x; r, t) = \frac{\sqrt{(x-x_L(x; r, t))(x_R(x; r, t)-x)}}{2\pi r x} \mathbf{1}_{(x_L(r, t), x_R(r, t))}(x), \quad x_L(r, t) := (1 - \sqrt{r})^2 t, \quad x_R(r, t) := (1 + \sqrt{r})^2 t. \quad (1.7)$$

## 2. Three-parametric Marcenko-Pastur density

PDE(1.5), with starting from  $a \geq 0$ , the Green function, is obtained by the solution of equation, [2]

$$z = \frac{1}{G_{\delta_0}(z)} + \frac{t}{1 - r t G_{\delta_0}(z)} + \frac{a}{(1 - r t G_{\delta_0}(z))^2}, \quad z \in \mathbb{C} \setminus \mathbb{R}, \quad r \in (0, 1], \quad t \geq 0, \quad a \geq 0.$$

We write the probability density via the Sokhotski-Plemelj theorem as  $\rho_{\delta_a}$  and call it the **three-parametric Marcenko-Pastur density**.  $\rho_{\delta_a}$  is given by the following explicit formula  $\rho(x; r, t, a)$ . In Figure 1, the original MP density and the three-parametric MP density are represented by curves. These histograms are generated from finite random matrices with corresponding parameters.

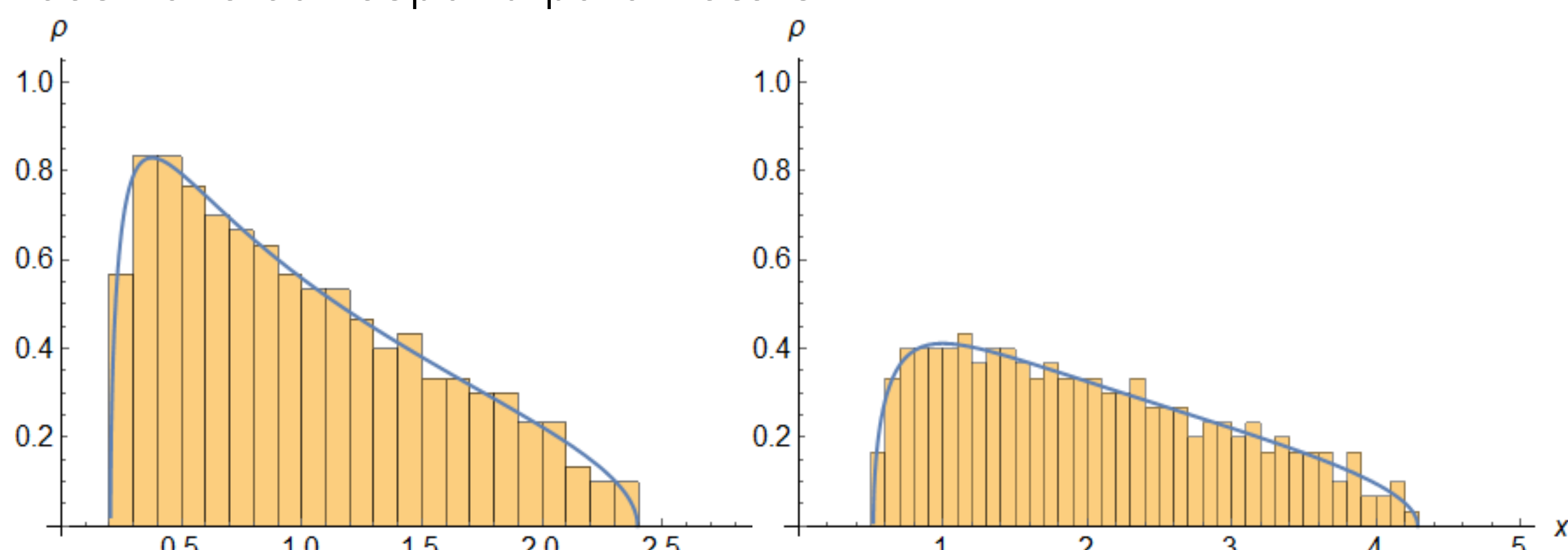


Figure 1: Histograms of eigenvalues of matrices  $L$  given by random rectangular matrices  $K$  with size  $1000 \times 300$ .

$K$ 's elements are randomly generalized following complex normal distribution, with mean 0 (right), and  $\sqrt{300}\delta_{ij}$  (left). MP density are shown by curves.

## 3. Main Theorem

The **three-parametric Marcenko-Pastur density**  $\rho(x; r, t, a)$  is given by,

$$\rho(x; r, t, a) = \frac{\sqrt{(x-f_L(x; r, t, a))(f_R(x; r, t, a)-x)}}{2\pi r x t} \mathbf{1}_{(x_L(r, t, a), x_R(r, t, a))}(x),$$

with,

$$f_L(x; r, t, a) = \left( \sqrt{\frac{t-a+\sqrt{-4a\varphi+(t-a)^2}}{2}} - \sqrt{\varphi+x+\frac{t-a+\sqrt{-4a\varphi+(t-a)^2}}{2}} \right)^2,$$

$$f_R(x; r, t, a) = \left( \sqrt{\frac{t-a+\sqrt{-4a\varphi+(t-a)^2}}{2}} + \sqrt{\varphi+x+\frac{t-a+\sqrt{-4a\varphi+(t-a)^2}}{2}} \right)^2,$$

where,

$$\varphi(x; r, t, a) := -\frac{2}{3}\{x - (r-1)t\} - \frac{\frac{2}{3}x^2 + \{3a - (2r+1)t\}x + t^2(r-1)^2 - \frac{g^3}{3 \times 2^{3/2}}}{g^3},$$

$$g(x; r, t, a) := -2x^3 + 3\{(2r+1)t + 6a\}x^2 - 3\{(r-1)\{(2r+1)t - 3a\}t - \sqrt{-3S}\}x + 2(r-1)^3 t^3,$$

$$S(x; r, t, a) := 4ax^3 - \{8a^2 + 4a(3r+2)t - t^2\}x^2 + 2[2a^3 - 2a^2(5r-2)t + a\{r(6r-1) + 1\}t^2 - (r+1)t^3]x + (r-1)^2 t^2 \{a^2 - a(4r-2)t + t^2\}.$$

$x_L(r, t, a)$ ,  $x_R(r, t, a)$ : Let  $x_1, x_2, x_3$  are real solutions of  $S(x; r, t, a) = 0$ , and  $x_1 \leq x_2 \leq x_3$ . Define  $x_L(r, t, a) := x_2$ ,  $x_R(r, t, a) := x_3$ .

## 5. References

- [1] Blaizot, J.-P., Nowak, M. A., Warchol, P.: Universal shocks in the Wishart random-matrix ensemble. Phys. Rev. E **87**, 052134/1–10 (2013)
- [2] Blaizot, J.-P., Nowak, M. A., Warchol, P.: Universal shocks in the Wishart random-matrix ensemble. II. Nontrivial initial conditions. Phys. Rev. E **89**, 042130/1–7 (2014)
- [3] Endo, T., Katori, M.: Three-Parametric Marcenko-Pastur Density. arXiv: math. PR/1907.07413

## 4. Proposition

(1) If and only if  $r = 1$ , the domain  $\{(x_L(r, t, a), x_R(r, t, a)) : t \geq 0\}$  touches the  $x = 0$ .

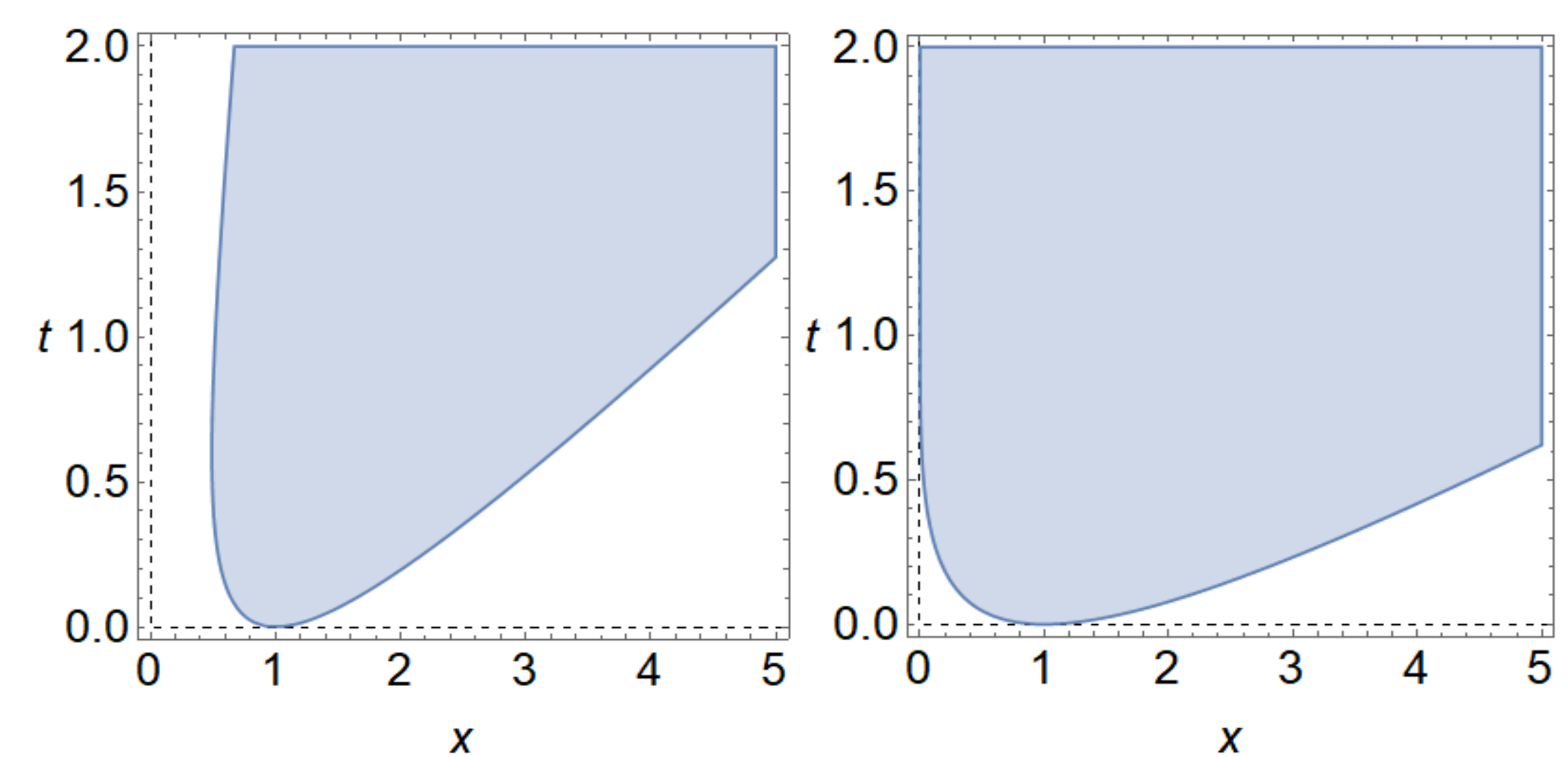


Figure 2: Time evolution of the support of the three-parametric Marcenko-Pastur Density starting from 1. Left figure:  $r = 0.3$ , Right figure:  $r = 1$ .

(2) There is a critical time  $t_c(a) = a$  such that time domain touches  $x = 0$ . More specifically, while  $0 \leq t \leq t_c(a)$ ,

$$x_L(r, t, a) \simeq \frac{4}{27a^2} (t_c(a) - t)^3, \text{ as } t \nearrow t_c(a)$$

(3) When, the three-parametric Marcenko-Pastur density shows the following dynamic critical phenomena at  $t = t_c(a)$ .

(i) For  $0 < t < t_c(a)$ ,

$$\rho(x; 1, t, a) \simeq \frac{9a}{4\pi} (t_c(a) - t)^{-5/2} (x - x_L(1, t, a))^{1/2}$$

as  $x \searrow x_L(1, t, a)$ ,  $t \nearrow t_c(a)$ .

(ii) For  $t = t_c(a)$ ,

$$\rho(x; 1, t_c(a), a) \simeq \frac{\sqrt{3}}{2\pi} a^{-2/3} x^{-1/3}$$

as  $x \searrow 0$ .

(iii) For  $t_c(a) < t$ ,

$$\rho(x; 1, t, a) \simeq \frac{1}{\pi t_c(a)} (t - t_c(a))^{1/2} x^{-1/2}$$

as  $x \searrow 0$ ,  $t \nearrow t_c(a)$ .

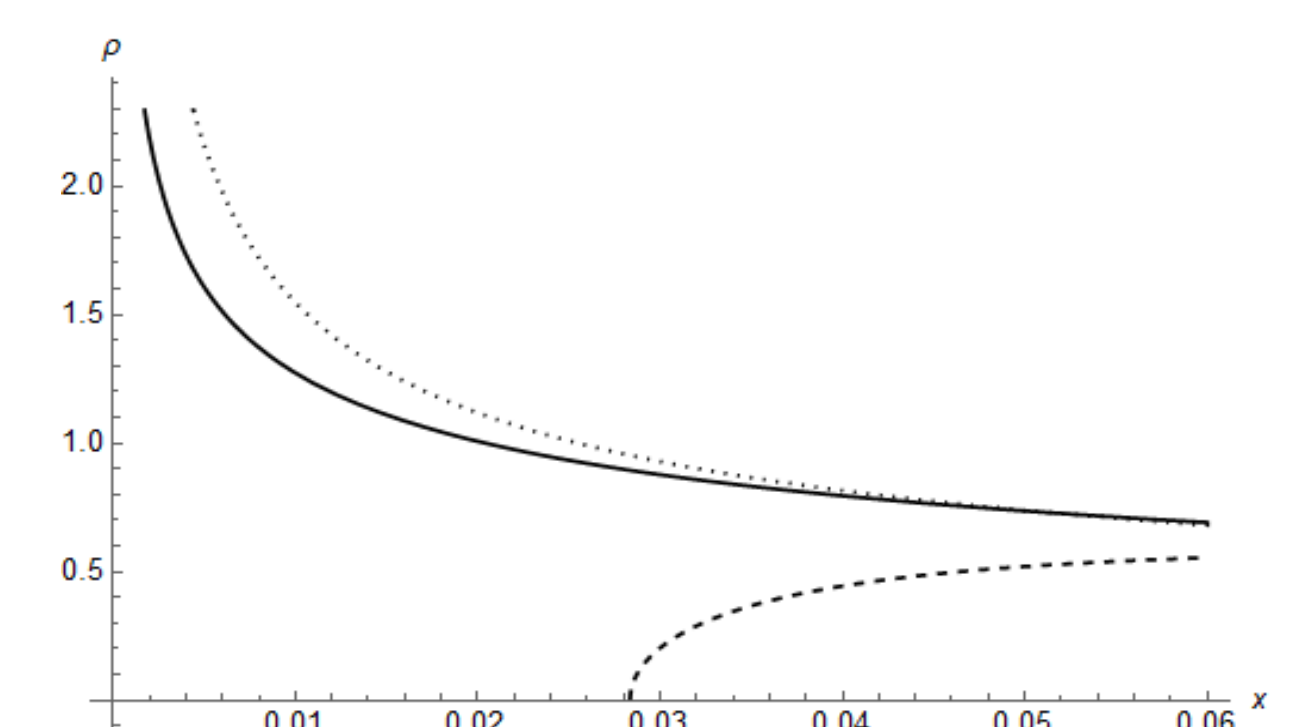


Figure 3: For (3) in  $r = 1$ ,  $a = 1$ . The dashed curve: (i) at  $t = 0.5t_c(1)$ . The solid curve: (ii) at the critical time  $t = t_c(1)$ . The dotted curve: (iii) at  $t = 1.5t_c(1)$ .

In Figure 3, the dashed curve denotes the emergence of  $\rho$  at  $x = x_L > 0$  with the critical exponent  $1/2$  at subcritical time. The solid curve  $\sim x^{-1/3}$  is weaker than the dotted curve  $x^{-1/2}$ .